Math 11 Calculus-Based Introductory Probability and Statistics

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Today:

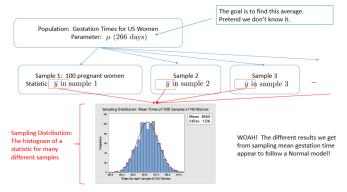
- Hypothesis Testing for Means (t-test)
- Better Understanding Errors Made in Hypothesis Tests

Inference for Means

We want to do all the same inference for means that we did with proportions: understanding the sampling distribution, hypothesis testing, and confidence intervals

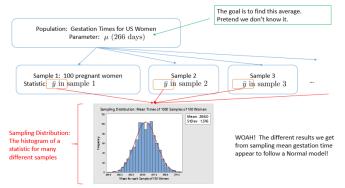
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In a previous class, we learned that the sampling distribution is approximately Normal with

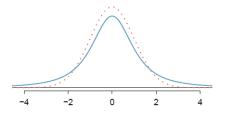
$$\mu_{model} = \mu_{population}$$
 and $\sigma_{mode} = \frac{\sigma_{model}}{\sqrt{n}}$.

Time for the Truth...

With smaller sample sizes (often n < 30) or populations where you don't know σ (and must approximate it using s_x), there is a better approximation of the sampling distribution than the Normal model.

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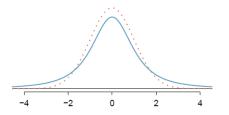


Solid line: t-distribution Dotted line: Normal (z) distribution Like the Normal distribution, the *t*-distribution is unimodal and symmetric.

Note, however, that the tails of the *t*-distribution are thicker, and this can change the area under the curve, and hence, the *p*-values that are based on it.

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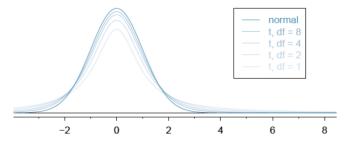
Note, however, that the tails of the *t*-distribution are thicker, and this can change the area under the curve, and hence, the *p*-values that are based on it.

When Should I Use Z vs. T?

- If you know σ (almost never true): Use a z-distribution.
- In all other cases: Use a t-distribution.

From z to t

In practice, we tend **not** to use the Normal curve as the approximation of the sampling distribution because the *t*-distribution gives us more precise results.



There are many curves in the *t*-distribution family. You choose the appropriate one based on the size of your sample.

With n data points in the sample, you use the t-distribution with df = n - 1.

df stands for "degrees of freedom" and as it gets bigger, the *t*-distribution morphs into a standard normal distribution.

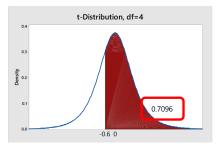
How much area is to the right of -0.6 on t_4 ?

As before: Graph >> Probability Distribution Plot >> View Probability

Probability Distribution Plot: View Probability				
Distribution Shaded Area				
	Distribution:			
	Degrees of freedom: 4			

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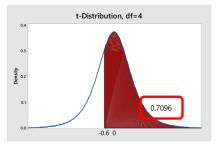
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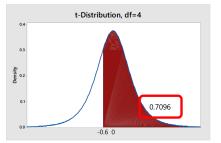
What percent of the t_6 distribution is more than 2 unites from 0?

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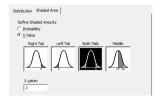
As before: Graph >> Probability Distribution Plot >> View Probability **B** 1 1 1 1 1

Double to the second

Probability Distribution Plot: View Probability					
Distribution Shaded Area					
	Distribution:				
	Degrees of freedom: 4				



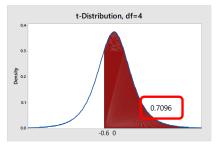
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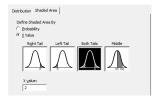
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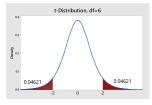
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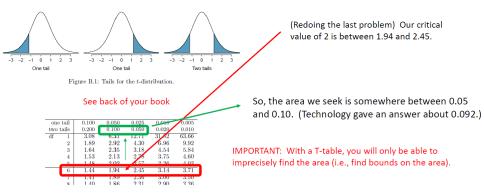


 $2 \times 0.046 = 0.092.$ (On N(0, 1), this answer would be about 1 - 0.95 = 0.05) 5/38

Using Tables with t-Distributions

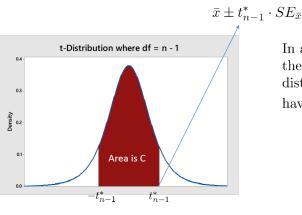
Since they don't want to print a new table for every possible df value, they print one table, but it is not as good as a z-table (standard Normal).

Also, a t-table is usually reversed (areas on outside, critical values in the table).



New, Improved CIs for Means!

For means, our CIs will use the same setup as before, but we live on a t-distribution, not a z-distribution:



In a sample with n data points, the best model for the sampling distribution is t_{n-1} . We still have $SE_{\bar{x}} = \frac{\sigma}{\sqrt{n}} \approx \frac{s_x}{\sqrt{n}}$

On average, how much do U.S. baby girls weigh? To find out, you sneak into random hospitals and collect illegally the weight of 12 random newborn babies. If $\bar{x} = 7.3$ lbs and $s_x = 2$, find a 90% C.I. for μ , the average weight of all U.S. female babies.

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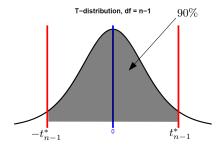
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Since n = 12, our sampling distribution is modeled by $t_{n-1} = t_{11}$.

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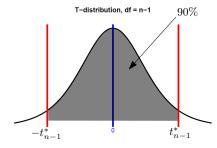
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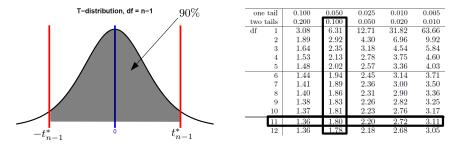


one t	ail	0.100	0.050	0.025	0.010	0.005
two ta	ils	0.200	0.100	0.050	0.020	0.010
df	1	3.08	6.31	12.71	31.82	63.66
	2	1.89	2.92	4.30	6.96	9.92
	3	1.64	2.35	3.18	4.54	5.84
	4	1.53	2.13	2.78	3.75	4.60
	5	1.48	2.02	2.57	3.36	4.03
	6	1.44	1.94	2.45	3.14	3.71
	7	1.41	1.89	2.36	3.00	3.50
	8	1.40	1.86	2.31	2.90	3.36
	9	1.38	1.83	2.26	2.82	3.25
	10	1.37	1.81	2.23	2.76	3.17
	11	1.36	1.80	2.20	2.72	3.11
-	12	1.36	1.78	2.18	2.68	3.05

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1) Decide on the model and find the critical value for your confidence level.

Since n = 12, our sampling distribution is modeled by $t_{n-1} = t_{11}$.



We read $t^* = 1.80$.

2) Break out the C.I. formula.

$$\bar{x} \pm t_{n-1}^* \times SE_{\bar{x}} \simeq 7.3 \pm 1.8 \times \frac{2}{\sqrt{12}}$$

= (6.26, 8.34).

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3) Interpret in plain English.

We are 90% confident that μ is between 6.26 and 8.34 pounds.

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$$\bar{x} \pm t_{n-1}^* \times SE_{\bar{x}} \simeq 7.3 \pm 1.8 \times \frac{2}{\sqrt{12}}$$

= (6.26, 8.34).

3) Interpret in plain English.

We are 90% confident that μ is between 6.26 and 8.34 pounds.

Remark: The actual value for μ is 7.5 pounds. Our C.I. was one of the lucky 90%.

You decide to build an 80% C.I. for some mean you are estimating. What is the critical value if your sample has size 15?

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df = 1	3.08	6.31	12.71	31.82	63.66
2	1.89	2.92	4.30	6.96	9.92
3	1.64	2.35	3.18	4.54	5.84
4	1.53	2.13	2.78	3.75	4.60
5	1.48	2.02	2.57	3.36	4.03
6	1.44	1.94	2.45	3.14	3.71
7	1.41	1.89	2.36	3.00	3.50
8	1.40	1.86	2.31	2.90	3.36
9	1.38	1.83	2.26	2.82	3.25
10	1.37	1.81	2.23	2.76	3.17
11	1.36	1.80	2.20	2.72	3.11
12	1.36	1.78	2.18	2.68	3.05
13	1.35	1.77	2.16	2.65	3.01
14	1.35	1.76	2.14	2.62	2.98
15	1.34	1.75	2.13	2.60	2.95
16	1.34	1.75	2.12	2.58	2.92
17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
21	1.32	1.72	2.08	2.52	2.83
22	1.32	1.72	2.07	2.51	2.82
23	1.32	1.71	2.07	2.50	2.81

• 1.34

- 1.35
- 1.75
- 1.76

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17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
21	1.32	1.72	2.08	2.52	2.83
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• 1.34

• 1.76

Answer: $t_{n-1}^* = t_{14}^* = 1.35$.

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About how much area is to the left of -2 on t_{15} ?

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df 1	3.08	6.31	12.71	31.82	63.66
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16	1.34	1.75	2.12	2.58	2.92
17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
21	1.32	1.72	2.08	2.52	2.83
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- $\bullet\,$ Between 0.1 and 0.2
- $\bullet\,$ Between 0.05 and 0.1
- Between 0.025 and 0.05
- Between 0.001 and 0.025

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- $\bullet\,$ Between 0.1 and 0.2
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- Between 0.025 and 0.05
- Between 0.001 and 0.025

Answer: Between 0.025 and 0.05

Suppose we want to do inference on a mean using a t-distribution. When does a t-distribution actually model the sample distribution?

- 1. It always does. It doesn't matter what the sample size is or what the population distribution looks like.
- 2. We only need the data points in our sample to be independent
- 3. If the data in the sample were chosen at random and are ${<}10\%$ of the population
- 4. We need randomization, <10 % population, and the population to be reasonably normal looking (skew is ok with a larger n).

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Answer: 4.

Hypothesis Testing with t-Distributions

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Let μ be the average time of finishers in the current year.

$$H_0: \mu = 3.683; H_A: \mu \neq 3.683$$

Hypothesis Testing with t-Distributions

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Let μ be the average time of finishers in the current year.

$$H_0: \mu = 3.683; H_A: \mu \neq 3.683$$

You collect data for 20 random runners in the current year, and get $\bar{x} = 3.8$ with $s_x = 0.5$. Run a hypothesis test with $\alpha = 0.05$.

You should **always** convert to the standard t-distribution. To do so, you'll need the t-score:

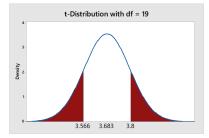
$$T = \frac{\text{point estimate} - \text{null value}}{SE}$$
$$= \frac{3.8 - 3.683}{0.5/\sqrt{20}} \simeq 1.046.$$

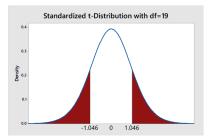
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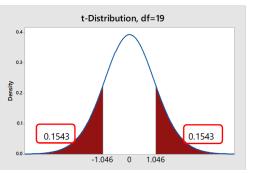
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$$= \frac{3.8 - 3.683}{0.5/\sqrt{20}} \simeq 1.046.$$











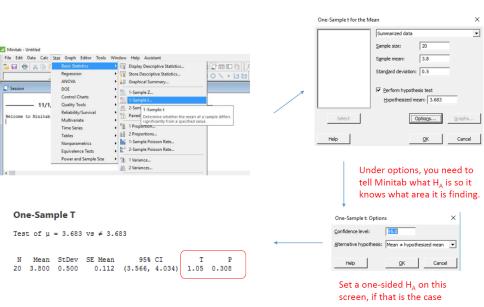
In a table, our T-score of 1.046 is to the left of 1.33. So we have more area than the two-tails value of 0.2. So, P-value > 0.2 > 0.05, as before.

Technology says the area is 0.154 + 0.154 = 0.308

Since P = 0.31 > 0.05, we do not reject H_0 . Our data are not strong enough to show a difference (if there is one).

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df 1	3.08	6.31	12.71	31.82	63.66
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5	1.48	2.02	2.57	3.36	4.03
6	1.44	1.94	2.45	3.14	3.71
7	1.41	1.89	2.36	3.00	3.50
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15	1.34	1.75	2.13	2.60	2.95
16	1.34	1.75	2.12	2.58	2.92
17	1.33	1.74	2.11	2.57	2.90
18	1 2 2	1.72	2.10	9.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
01	1.90	1.79	9.09	0 50	9.09

The Last Example Done Easily



What Assumptions Must We Meet to Use *t*-Distributions?

What Assumptions Must We Meet to Use *t*-Distributions?

When doing statistical inference (using CIs or hypothesis testing), you rely on the shape and SE of the sampling distribution. To get these to be correct, you must have:

Independence of data: Knowing one piece of data should not help you predict other data. Since this is hard to check, this is usually replaced with the:

- Randomization Condition
- < 10% Condition

Population distribution must be nearly normal. Since this is hard to check:

- Look for near normality in the histogram of your sample (or its qq-plot).
- More skew is OK as n gets larger.

We study beetle biodiversity in a pasture. For this, we collect a biodiversity index (Steinhaus index) in 12 parcels and get the following data:

 $\bar{x} = 0.2505$ and $s_x = 0.0959$.

We study beetle biodiversity in a pasture. For this, we collect a biodiversity index (Steinhaus index) in 12 parcels and get the following data:

 $\bar{x} = 0.2505$ and $s_x = 0.0959$.

An environmental engineer, specialist of beetle populations, tells you than an average biodiversity index lower than 0.49 in the pasture would be worrying.

Build a test with level of confidence $\alpha = 5\%$ for determining if the state of the pasture is worrying.

1) We build hypotheses for the situation

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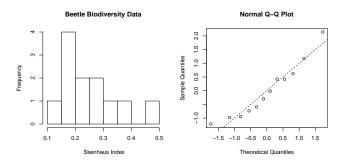
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3) t-score your data

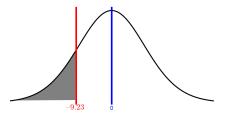
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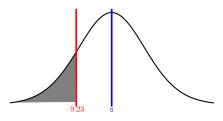


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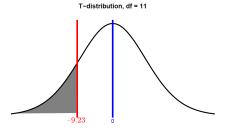


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With Minitab, you find the area under the curve is $\simeq 8 \times 10^{-7}$.

Since $p \simeq 8 \cdot 10^{-7} \ll 0.05$, we reject H_0 and favor H_A .

There is (a very) strong evidence that the true average biodiversity index is smaller than 0.49.

There are Four possible scenarios we must think about related to hypothesis testing:

- We reject H_0 when H_0 is actually true
- We do not reject H_0 when H_0 is actually true
- We reject H_0 when H_0 is actually false
- We do not reject H_0 when H_0 is actually false

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Type I Error

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Type I Error Awesome! Awesome! Type II Error

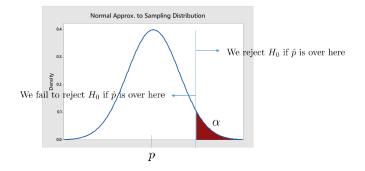
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٠	We reject H_0 when H_0 is actually true	Type I Error
•	We do not reject H_0 when H_0 is actually true	Awesome!
•	We reject H_0 when H_0 is actually false	Awesome!
•	We do not reject H_0 when H_0 is actually false	Type II Error

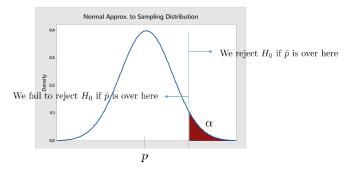
Test Conclusion Truth	Do not reject H_0	Reject H_0
H_0 true	OK	Type I Error
H_A true	Type II Error	OK

 $P(\text{Type I error}) = P(\text{reject } H_0 | H_0 \text{ is true}).$

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 $P(\text{Type I error}) = \alpha$

We reject the null when the *p*-value is $< \alpha$ (say, 0.05). For this to happen, \hat{p} must be large enough to get the shaded area $< \alpha$. Said differently, we need \hat{p} to be in the top 5% of the data values. This happens with probability α .

We also care about

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The **power** of a test is

P (Reject $H_0|H_0$ is false) = $1 - \beta$.

You'll learn how to deal with β in more advanced stat classes.

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Moral: For most actions,

Type I Error \searrow when Type II Error \nearrow (and vice versa).

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What does Type I Error mean in this case?

A Type I Error would mean the restaurant did meet regulations, but the inspector thought this was not true and shut it down.

A Type II Error would mean the restaurant violated regulations, but the inspector made the mistake of letting the restaurant stay open.

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Which error is worse for the restaurant owner? Which is worse for the diners?

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Which error is worse for the restaurant owner? Which is worse for the diners?

The restaurant owner dislikes Type I because the restaurant is shut down for no reason. Diners dislike Type II because they are exposed to dangerous conditions.

A student claims she can guess the suit on the top card from a shuffled deck better than chance. You decide to design an experiment to test this.

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If the student has no powers, she will get the correct answer one-fourth of the time. We set $H_0 p = 0.25$.

Since she claims to do better, we have H_A : p > 0.25.

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We eventually look at a sampling distribution, which will be normal if we have at least 10 successes and failures. We expect $0.25 \times n$ successes, so we need at least $n \geq 40$ to get adequate successes.

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7. We decide on a sample of size 100 because $100 \ge 40$ and drawing a larger sample isn't too hard. During the experiment, she guesses 32 suits correctly. She says this proves her claim. Why is her logic false?

We got $\hat{p} = 0.32$, but sampling variability exists. Some samples of 100 cards would yield results like this, other may not support her claim. We must sort out whether 32% is from sampling variation or not.

Reshuffling gives independence, 100 cards i less than 10% of all the cards we could ask her $(= \infty!)$, and the sample is big enough to expect at least 10 successes and failures.

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Under the null, $\mu(\hat{p}) = p = 0.25$ and $SE(\hat{p}) = \sqrt{pq/n} = \sqrt{0.75 \times 0.25/100} \simeq 0.0433.$

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It is the probability that sample variation gives us the result $\hat{p} = 0.32$ (or something more extreme) if the null hypothesis is true (that is, if the student is just randomly guessing).

11. Using symbols, how can we fill in this blank for our particular problem?

$$p$$
-value = $P($)

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$$p$$
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- 12. Describe two ways to find this *p*-value.
 - We can use Minitab to find the area on the sampling distribution itself,
 - or we can convert \hat{p} to a z-score and use a table to find the area on the standard normal distribution.

Let's learn a new way to do this!

Minitab - Untitled

	Regression ANOVA	s Window Help Assistant Display Descriptive State Store Descriptive State All Graphical Summary		॓॓॓͡ □ □ ि □ □ □ [] /×	(2= -: A 70)	24		Minitab a tests and
Session Test and C Test of p = Sample X 1 32 Using the n	0.25 vi Multivariate N Si 100 0 Tables Nonparametrics	1. Sample Z 1. Sample L 2. Sample L 1. Paired L 2. Proportion 2. Proportion 2. Proportion 2. Sample L 1. Paired L 2. Proportion 2. Sample L 1. Proportion 2. Sample L 1. Proportion 2. Sample L	whether the proportion	of an event observed a specified value.				(without showing a thinking b
 < Worksheet 1 + C1	Power and Sample Size	① 1 Variance ② 2 Variances 11 Correlation 0 ² Covariance C5 ▲		C10 C11	C12 C13	C14	C15 C16 4	
One-Sample	Proportion Summarized data Nugber of events: Nughber of trials:	× 32 100	One-Sample P Confidence level Alternative hypo		n > hypothesized pro	×	Test and CI for Test of p = 0.	
Selec	E Perform hypothe		Method:	Normal ap	Oproximation	Cancel	1 32 10	N Sample p 95% Lo 0 0.320000 Wal approximation.

Minitab allows us to run tests and gives P-values (without drawing areas or showing any of the visual thinking behind the test).

rest and ci for one Proportion												
Test of $p = 0.25$ vs $p > 0.25$												
Sample 1	Х 32	N 100	Sample p 0.320000	95% Lower 0.	Bound 243272	Z-Value 1.62	P-Value 0.053)				

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14. This suggests that if we ran the experiment 20 times, one of those times we would expect to get a result like 0.32 (or wilder!). If there were 20 science labs in the country running this experiment, what would happen?

We expect one of them to get a result whose probability is very near the 0.05 significance threshold. They will probably publish a paper saying they found a student with special powers. If we live in a society that only publishes results with *p*-values < 0.05, then for every published finding, we might expect 19 unpublished

findings that found no interesting result.

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16. Give some reasons why a 0.05 cutoff is silly.

0.05 is a historical artifact derived from a 1931 book by Fisher (*The Design of Experiments*). He though a 1 in 20 event might be surprising enough to toss out one's belief system (H_0) . In some situations, you might be willing to reject the null hypothesis

In some situations, you might be willing to reject the null hypothesis more or less readily.

17. The probability that the null hypothesis is true is 5.3%.

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False. The cutoff, or α level, is decided by the researcher. Values of $\alpha = 0.05, 0.01, 0.001$ are common.

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21. If $\alpha = 0.10$, the p-value of 0.09 is statistically significant.

True. The phrase "statistically significant" means that the p-value is less than the α cutoff.

Note: Statistical significance is not the same as practical significance. For example, consider an SAT class that raises grades by 0.1% vs. a drug that cures 0.1% more people.

Your friend has two coins, one that is fair, and the other that comes up tails 85% of the time. You get to spin one of the coins and then must decide which coin it is.

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Set H_0 : p = 0.85 and H_A : p = 0.5.

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Find a p-value! It equals

 $P(\text{our result or something more extreme}|H_0 \text{ is true}).$

Here, this means $P(\text{Heads}|H_0) = 0.15$. We might reject H_0 .

Here, the *p*-value is $P(\text{Tails}|H_0) = 0.0.85$. We have a piece of information that meshes well with H_0 , so we probably won't reject it.

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4. What do Type I and Type II Errors mean in context?

	Heads Appears	Tails Appears
H_0 is true (unfair coin was flipped)	We think our coin is FAIR Type I Error	We think our coin is UNFAIR
H_0 is false (fair coin was flipped)	We think our coin is FAIR.	We think our coin is UNFAIR. Type 2 Error